

$\Rightarrow X, Y$ are independent

$$(1) E[X \cdot Y] = E[X] \cdot E[Y]$$

Exercise Prove this (easy)

$$(2) E[g(X) \cdot h(Y)] = E[g(X)] E[h(Y)]$$

Exercise

Prove that X, Y indep. implies $g(X), h(Y)$ are indep. (hard)

Hint: adapt proof in discrete case

#44, ch. 2, P. 134.

$$(3) \text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$$

Exercise Prove this (easy)

3.6 Bayes' Rule

original version :

$$P(A|B) = \frac{P(B|A) \cdot P(A)}{P(B)}$$

If $A_1, A_2, \dots, A_n$ form a Partition of Ω , by total probability

$$P(B) = \sum_{i=1}^n P(A_i \cap B)$$

$$= \sum_{i=1}^n P(B|A_i) \cdot P(A_i)$$

so for any A_k ($k=1, \dots, n$)

$$P(A_k|B) = \frac{P(B|A_k) \cdot P(A_k)}{\sum_{i=1}^n P(B|A_i) \cdot P(A_i)}$$

III If X, Y are discrete r.v.s

$$P_{X|Y}(x|y) = \frac{P_{Y|X}(y|x) \cdot P_X(x)}{\sum_t P_{Y|X}(y|t) \cdot P_X(t)}$$

IV If X, Y are cont. r.v.s

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) \cdot f_X(x)}{\int_{-\infty}^{\infty} f_{Y|X}(y|t) f_X(t) dt}$$

Ex.

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Let Y be normal with $\sigma=1$,
and $\mu=X$, another r.v. .

Suppose X is continuous uniform
on $[1, 3]$.

a.) find $f_Y(y)$

We are given

$$f_X(x) = \begin{cases} \frac{1}{2} & \text{if } 1 \leq x \leq 3 \\ 0 & \text{otherwise} \end{cases}$$

and

$$f_{Y|X}(y|x) = \begin{cases} \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{(y-x)^2}{2}} & 1 \leq x \leq 3 \\ & (y \in \mathbb{R}) \\ 0 & \text{otherwise} \\ & (y \in \mathbb{R}) \end{cases}$$

so

$$f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$

$$= \int_{-\infty}^{\infty} f_{Y|X}(y|x) \cdot f_X(x) dx$$

$$= \int_1^3 \frac{1}{2} \cdot \frac{e^{-\frac{(x-y)^2}{2}}}{\sqrt{2\pi}} dx$$

Sub.

$$u = x - y$$

$$du = dx$$

$$x=1 \rightarrow u=1-y$$

$$x=3 \rightarrow u=3-y$$

$$= \frac{1}{2} \int_{1-\gamma}^{3-\gamma} \frac{e^{-u^2/2}}{\sqrt{2\pi}} du$$

$$= \boxed{\frac{1}{2} (\Phi(3-\gamma) - \Phi(1-\gamma))}$$

b.) find $f_{X|Y}(x|\gamma)$.

$$f_{X|Y}(x|\gamma) = \frac{f_{Y|X}(\gamma|x) \cdot f_X(x)}{f_Y(\gamma)}$$

$$= \frac{\cancel{\frac{1}{2}} \cdot \frac{1}{\sqrt{2\pi}} \cdot e^{-(\gamma-x)^2/2}}{\cancel{\frac{1}{2}} (\Phi(3-\gamma) - \Phi(1-\gamma))}$$

$$= \begin{cases} \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{(y-x)^2}{2}}}{\Phi(3-y) - \Phi(1-y)} & 1 \leq x \leq 3 \\ 0 & \text{otherwise} \end{cases} \quad y \in \mathbb{R}$$

e.) Suppose we sample Y and get $y=3$. What is Probability that $X \leq 2$, i.e. find

$$P(X \leq 2 | Y=3)$$

(exercise) ans = .2848

d.) find $E[Y]$

$$\text{Ans.} = 2$$

4.1 Derived Distributions

$\Rightarrow$ If $Y = g(X)$, how to compute

$$f_Y(y)$$

from $f_X(x)$. ? We call

$f_X(x)$ a derived distribution.

2-Step Process

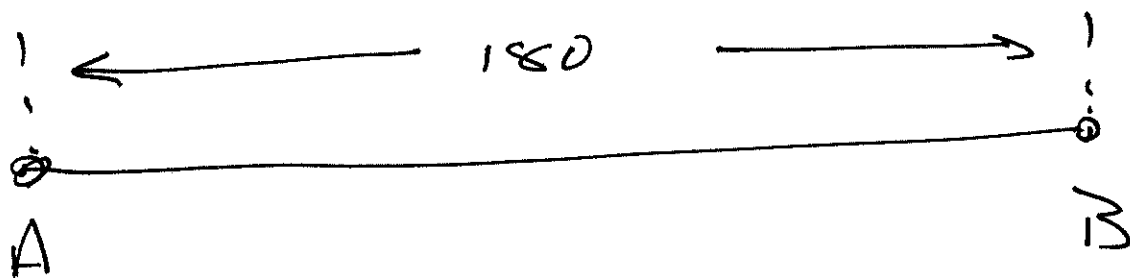
1.) find CDF

$$F_Y(y) = P(g(X) \leq y) = \int_{\{x | g(x) \leq y\}} f_X(x) dx$$

2.) differentiate

$$f_y(v) = \frac{d}{dy} \overline{F}_y(v)$$

EX A car travels a dist. of 180 miles at const. rate R , where R is continuous uniform on $30 \leq v \leq 60$ (mph)



let T be time at travel. find

$$f_T(t)$$

solution

From $180 = R \cdot T$, we get

$$T = \frac{180}{R}$$

$$\textcircled{1} F_T(t) = P(T \leq t)$$

$$= P\left(\frac{180}{R} \leq t\right)$$

$$= P\left(R \geq \frac{180}{t}\right)$$

$$= 1 - P\left(R < \frac{180}{t}\right)$$

$$= 1 - F_R\left(\frac{180}{t}\right)$$

(2) differentiate

$$f_T(t) = + f_R\left(\frac{180}{t}\right) \left(+ \frac{180}{t^2}\right)$$

We have

$$f_R(r) = \begin{cases} \frac{1}{30} & 30 \leq r \leq 60 \\ 0 & \text{otherwise} \end{cases}$$

Hence

$$f_T(t) = \begin{cases} \frac{1}{30} \cdot \frac{180}{t^2} & 30 \leq \frac{180}{t} \leq 60 \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{6}{t^2} & \text{if } 3 \leq t \leq 6 \\ 0 & \text{otherwise} \end{cases}$$

($a \neq 0$)
 If $Y = aX + b$, then

$$f_Y(y) = \frac{1}{|a|} \cdot f_X\left(\frac{y-b}{a}\right) \quad a \neq 0$$

Proof

$$\begin{aligned} \textcircled{1} \quad F_Y(y) &= P(Y \leq y) \\ &= P(aX + b \leq y) \\ &= \begin{cases} P\left(X \leq \frac{y-b}{a}\right) & a > 0 \\ P\left(X \geq \frac{y-b}{a}\right) & a < 0 \end{cases} \\ &= \begin{cases} P\left(X \leq \frac{y-b}{a}\right) & a > 0 \\ 1 - P\left(X < \frac{y-b}{a}\right) & a < 0 \end{cases} \end{aligned}$$

$$= \begin{cases} F_X\left(\frac{y-b}{a}\right) & a > 0 \\ 1 - F_X\left(\frac{y-b}{a}\right) & a < 0 \end{cases}$$

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$$(2) f_Y(y) = \begin{cases} f_X\left(\frac{y-b}{a}\right) \cdot \frac{1}{a} & a > 0 \\ -f_X\left(\frac{y-b}{a}\right) \cdot \frac{1}{a} & a < 0 \end{cases}$$

$$\therefore f_Y(y) = \frac{1}{|a|} \cdot f_X\left(\frac{y-b}{a}\right)$$

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Corollary

if X is normal with mean μ ,
variance σ^2 , then $Y = aX + b$
is also normal w. mean $a\mu + b$
and variance $a^2\sigma^2$, std. dev. $= |a| \cdot \sigma$

Defn

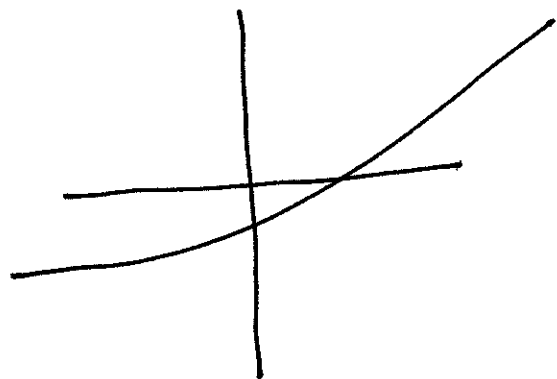
$g: \mathbb{R} \rightarrow \mathbb{R}$ is called strictly
monotonic iff either

$$(1) \quad x_1 < x_2 \rightarrow g(x_1) < g(x_2) \quad \left\{ \begin{array}{l} \text{strictly} \\ \text{increasing} \end{array} \right.$$

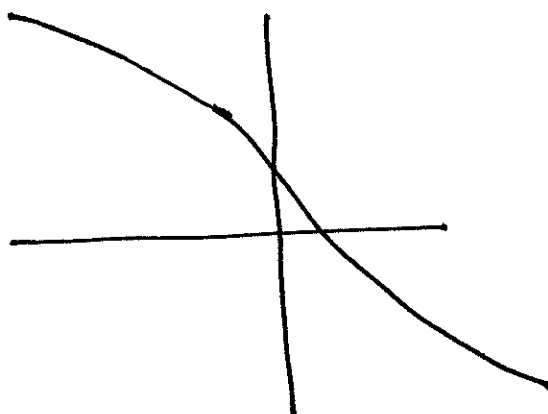
or

$$(2) \quad x_1 < x_2 \rightarrow g(x_1) > g(x_2) \quad \left\{ \begin{array}{l} \text{strictly} \\ \text{decreasing} \end{array} \right.$$

(1)



(2)



note: such functions are invertible

$$\text{let } h = g^{-1}$$

(1) $\Leftrightarrow h$ is strictly increasing

and

(2) $\Leftrightarrow h$ " " decreasing

If g is differentiable then

$$(1) \Leftrightarrow g' > 0 \Leftrightarrow h' > 0$$

$$(2) \Leftrightarrow g' < 0 \Leftrightarrow h' < 0$$

Theorem

let $y = g(x)$, then

$$f_Y(y) = f_X(h(y)) \cdot |h'(y)|$$

Proof

Suppose g is strictly increasing,
(leave decreasing as exercise).

Then

$$\begin{aligned}
 (1) \quad F_Y(y) &= P(Y \leq y) \\
 &= P(g(X) \leq y) \\
 &= P(X \leq h(y)) \quad \left\{ \begin{array}{l} \text{using} \\ h \text{ is} \\ \text{inc.} \end{array} \right. \\
 &= F_X(h(y))
 \end{aligned}$$

$$\begin{aligned}
 (2) \quad f_Y(y) &= f_X(h(y)) \cdot h'(y) \\
 &= f_X(h(y)) \cdot |h'(y)| \quad \left\{ \begin{array}{l} \text{since} \\ h' > 0 \end{array} \right.
 \end{aligned}$$



Ex. let $Y = X^2$ where X
is a cont. r.v. with

$$\text{range}(X) = [0, \infty)$$

Then also

$$\text{range}(Y) = [0, \infty)$$

so $g: [0, \infty) \rightarrow [0, \infty)$ is strictly
increasing, $g(x) = x^2$ has inverse
 $h(y) = \sqrt{y}$, Thus

$$f_Y(y) = \begin{cases} f_X(\sqrt{y}) \cdot \frac{1}{2\sqrt{y}} & (y \geq 0) \\ 0 & \text{otherwise} \end{cases}$$

$$\text{let } g: \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}$$

let X, Y be indep. r.v.s,
both ^{cont} uniform on $[0, 1]$

$$\text{let } \begin{array}{c} g \\ \downarrow \\ Z = \max(X, Y) \end{array}$$

We know

$$f_X(x) = \begin{cases} 1 & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$F_X(x) = \begin{cases} 0 & x < 0 \\ x & 0 \leq x \leq 1 \\ 1 & x > 1 \end{cases}$$

Similarly for Y .

goal : find f_Z

$$F_Z(z) = P(Z \leq z)$$

$$= P(\max(X, Y) \leq z)$$

$$= P(X \leq z, Y \leq z)$$

$$= P(X \leq z) \cdot P(Y \leq z)$$

$$= F_X(z) \cdot F_Y(z)$$

$$= z^2$$

$$\therefore f_Z(z) = \begin{cases} 2z & \text{if } 0 \leq z \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

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