

Supplemental Lecture

Total Probability Theorem

Let $A_1, \dots, A_n$ form a partition of Ω , and assume $P(A_i) > 0$ for $i=1, \dots, n$.

Let $B \subseteq \Omega$. Then

$$P(B) = P(A_1) \cdot P(B|A_1) + \dots + P(A_n) \cdot P(B|A_n)$$

Ex. Roll a 6-sided fair die. If the result $\in \{1, 2, 3\}$, roll again, then stop. If 1st roll $\in \{4, 5, 6\}$, no second roll.

$$\text{Sum} = \begin{cases} 1^{\text{st}} \text{ roll} + 2^{\text{nd}} \text{ roll} & \text{if } 1^{\text{st}} \in \{1, 2, 3\} \\ 1^{\text{st}} \text{ roll} & \text{if } 1^{\text{st}} \in \{4, 5, 6\} \end{cases}$$

Find $P(\text{sum} \geq 5)$.

Let $A_i = \{1^{\text{st}} \text{ roll is } i\}$ ($1 \leq i \leq 6$).

Then $P(A_i) = \frac{1}{6}$ ($1 \leq i \leq 6$). Let

$$B = \{\text{sum} \geq 5\}.$$

Note:

if A_1 , then $B \iff 2^{\text{nd}} \text{ roll} \in \{4, 5, 6\}$

if A_2 , then $B \iff 2^{\text{nd}} \text{ roll} \in \{3, 4, 5, 6\}$

if A_3 , then $B \iff 2^{\text{nd}} \text{ roll} \in \{2, 3, 4, 5, 6\}$

Thus

$$P(B|A_1) = \frac{3}{6} = \frac{1}{2}$$

$$P(B|A_2) = \frac{4}{6} = \frac{2}{3}$$

$$P(B|A_3) = \frac{5}{6}$$

$$P(B|A_4) = 0$$

$$P(B|A_5) = 1$$

$$P(B|A_6) = 1$$

Hence

$$\begin{aligned}
 P(B) &= P(B|A_1) \cdot P(A_1) + \dots + P(B|A_6) \cdot P(A_6) \\
 &= \frac{1}{2} \cdot \frac{1}{6} + \frac{2}{3} \cdot \frac{1}{6} + \frac{5}{6} \cdot \frac{1}{6} + 0 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} + 1 \cdot \frac{1}{6} \\
 &= \frac{1}{6} \left(\frac{1}{2} + \frac{2}{3} + \frac{5}{6} + 0 + 1 + 1 \right) \\
 &= \frac{1}{6} \left(\frac{24}{6} \right) = \frac{4}{6} = \boxed{\frac{2}{3}}
 \end{aligned}$$

Another way:

Ω {

11	12	13	14	15	16
21	22	23	24	25	26
31	32	33	34	35	36
4					
5					
6					

← $B = \{\text{sum} \geq 5\}$

$$P(B) = \frac{12}{36} + \frac{2}{6} = \frac{1}{3} + \frac{1}{3} = \boxed{\frac{2}{3}}$$

Bayes' Rule

Recall $P(A \cap B) = P(B) \cdot P(A|B)$

$$\parallel$$

$$P(B \cap A) = P(A) \cdot P(B|A)$$

$$\therefore P(B) \cdot P(A|B) = P(A) \cdot P(B|A)$$

$$\therefore \boxed{P(A|B) = \frac{P(A) \cdot P(B|A)}{P(B)}}$$

This simplest form of Bayes' rule

Let $A_1, \dots, A_n$ be a partition of Ω and $P(A_i) > 0$. Assume $P(B) > 0$.

Then for any $i = 1, \dots, n$, we have

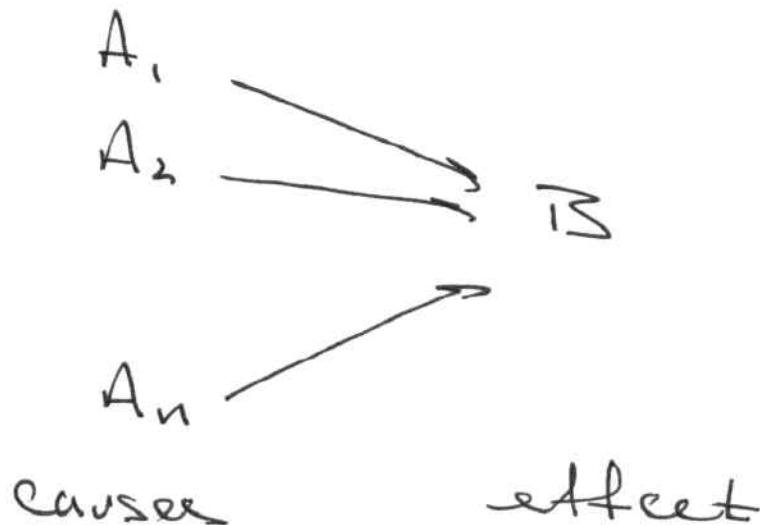
$$P(A_i|B) = \frac{P(A_i) \cdot P(B|A_i)}{P(A_1) \cdot P(B|A_1) + \dots + P(A_n) \cdot P(B|A_n)}$$

This is called Bayes's Rule.

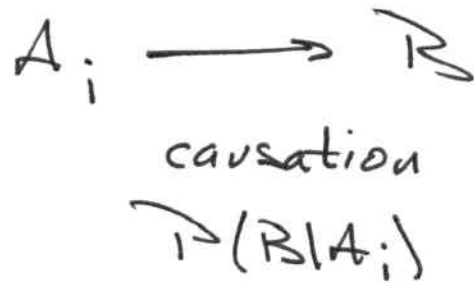
Bayes is used for 'inference' Problems.
We have causes:

$A_1, A_2, \dots, A_n$

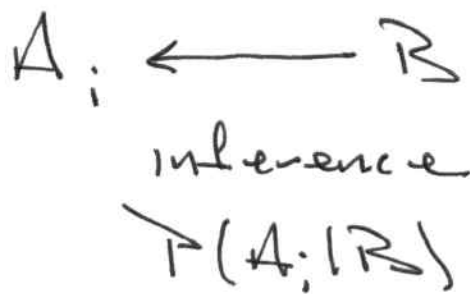
that may produce an effect B .



note $\rightarrow (R|A_i)$ models strength of causation



often, we wish to go from effect to cause. $\rightarrow (A_i|R)$ models the strength of the inference of A_i from R .



$$\underbrace{\rightarrow (A_i|R)}_{\text{Posterior}} = \frac{\underbrace{\rightarrow (A_i)}_{\text{Prior}} \underbrace{\rightarrow (R|A_i)}_{\text{Likelihood}}}{\underbrace{\rightarrow (R)}_{\text{marginal}}}$$

Ex. Disease example again

$T = \{ \text{test positive} \}$

$D = \{ \text{has disease} \}$

Given information:

$$P(T|D) = .99$$

$$P(T|D^c) = .10$$

$$P(D) = .05, \quad P(D^c) = .95$$

find $P(\text{disease} | \text{test})$, i.e.

$$P(D|T) = \frac{P(D) \cdot P(T|D)}{P(T)}$$

$$= \frac{P(D) \cdot P(T|D)}{P(D) \cdot P(T|D) + P(D^c) \cdot P(T|D^c)}$$

$$= \frac{(.05)(.99)}{(.05)(.99) + (.95)(.10)} = .34256 \approx \boxed{.3426}$$

round to 4 digits

hw1 #2

Players ranked: $1 < 2 < 3$

Assume
Play in
order
1, 2, 3

$$\left. \begin{aligned} A &= \{ \text{win against 1} \}, P(A) = P_1 \\ B &= \{ \text{" " 2} \}, P(B) = P_2 \\ C &= \{ \text{" " 3} \}, P(C) = P_3 \end{aligned} \right\} \quad P_1 > P_2 > P_3$$

$$\begin{aligned} E &= \{ \text{win tournament} \} \\ &= (A \cap B) \cup (B \cap C) \end{aligned}$$

need

assume $= P(B)$

$$P(A \cap B) = P(A) \cdot \overline{P(B|A)} = P(A) \cdot P(B) = P_1 \cdot P_2$$

$$P(B \cap C) = \dots \dots \dots = P_2 \cdot P_3$$

$$P(A \cap B \cap C) = \dots \dots \dots = P_1 \cdot P_2 \cdot P_3$$

1.5 Independence

It's possible that knowledge of $B \subseteq \Omega$ tells us nothing about $A \subseteq \Omega$. i.e.

$$* \quad P(A|B) = P(A)$$

In this case we say A is independent of B . Recall

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Thus * gives

**

$$P(A \cap B) = P(A) \cdot P(B)$$

Defn

events A, B are independent iff ** holds.

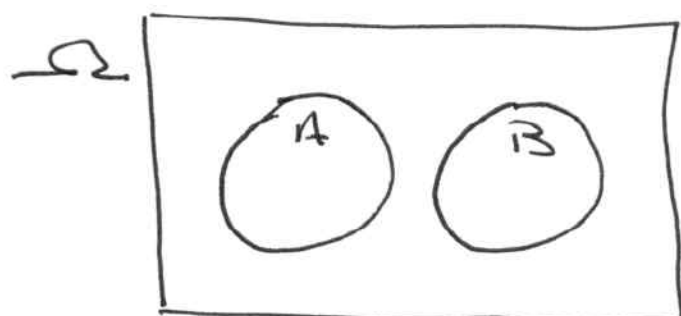
Independence has a simple intuition:

A, B represent physically non-interacting processes.

Independence is not easily visualized as a Venn diagram

note $A \cap B = \emptyset \not\Rightarrow A, B$ independent

why? $A \cap B = \emptyset \Rightarrow P(A \cap B) = 0$. But neither $P(A)$ nor $P(B)$ need be impossible



$$A \cap B = \emptyset$$

$$P(A) > 0$$

$$P(B) > 0$$

$$\therefore P(A) \cdot P(B) > 0$$

but

$$P(A \cap B) = 0$$

111

Ex. Throw 2 fair dice. let

$$A_i = \{1^{\text{st}} \text{ die is } i\} \quad 1 \leq i \leq 6$$

$$B_j = \{2^{\text{nd}} \text{ die is } j\} \quad 1 \leq j \leq 6$$

fair means $P(A_i) = \frac{1}{6}$, $P(B_j) = \frac{1}{6}$

Assume A_i, B_j are independent.

Then

$$P((i,j)) = P(A_i \cap B_j) = P(A_i) \cdot P(B_j) = \frac{1}{36}$$

which is the discrete uniform Probability law.

$$\text{Let } C = \{\text{sum} = 5\}$$

$$= \{14, 23, 32, 41\}$$

$$\therefore P(C) = \frac{4}{36} = \frac{1}{9}$$

Is C independent of A_2 ?

$$A_2 = \{21, 22, 23, 24, 25, 26\}$$

$$\therefore A_2 \cap C = \{23\}$$

$$\therefore P(A_2 \cap C) = \frac{1}{36}$$

But

$$P(A_2) \cdot P(C) = \frac{1}{6} \cdot \frac{1}{9} = \frac{1}{54} \neq \frac{1}{36}$$

$\therefore A_2, C$ are not independent.

$$\text{Let } D = \{\text{double}\} = \{11, 22, 33, 44, 55, 66\}$$

$$\therefore P(D) = \frac{6}{36} = \frac{1}{6}$$

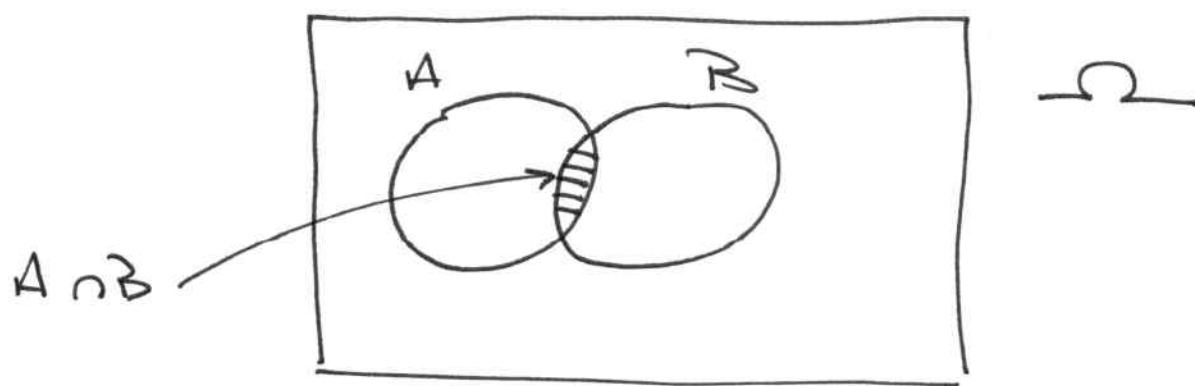
$$D \cap A_2 = \{22\}$$

$$\therefore P(D \cap A_2) = \frac{1}{36}$$

$$\text{But } P(D) \cdot P(A_2) = \frac{1}{6} \cdot \frac{1}{6} = \frac{1}{36} = P(D \cap A_2)$$

Thus A_2, D are independent. $\blacksquare$

Assume Probability is normalized area, as in a continuous model



$$P(A|B) = P(A)$$

$$\frac{P(A \cap B)}{P(B)} = P(A)$$

$$\frac{\frac{\text{area}(A \cap B)}{\cancel{\text{area}(\Omega)}}}{\frac{\text{area}(B)}{\cancel{\text{area}(\Omega)}}} = \frac{\text{area}(A)}{\text{area}(\Omega)}$$

assume
 $\text{area}(\Omega) = 1$

$$\therefore \frac{\text{area}(A \cap B)}{\text{area}(B)} = \text{area}(A)$$