

# Time Series Lecture Module 2

- Topics in this lecture
  - Stationarity and Unit Roots
  - Spurious Regressions
  - Cointegration
  - Error-Correction Models

# Stationarity and Unit Roots

- We need a precise test to distinguish between stationarity and non-stationarity
  - Mean is unknown and variance explodes for non-stationary time series
- Graphical techniques were not based on any precise statistical test
- In this set of slides, we'll discuss the "unit root", and how to identify it
  - Though a unit root has a precise definition, it basically summarizes when a autoregressive relationship is non-stationary
  - Generally, we take differences, or differences of differences, or differences of differences of...of differences of differences to purge an autoregressive relationship of non-stationary properties.

# What is a unit-root?

- Consider the following AR(1) model

$$y_t = \phi_1 y_{t-1} + u_t$$

- Three possible cases for this AR(1) model:

- 1  $|\phi_1| < 1$ , and the series is stationary
- 2  $\phi_1 = 1$ , and the series has a unit root and is non-stationary
- 3  $\phi_1 = -1$ , and the series is non-stationary without a unit root
- 4  $|\phi_1| > 1$ , and the series is explosive

- To test for unit root, first subtract  $y_{t-1}$  from both sides.

$$y_t - y_{t-1} = (\phi_1 - 1)y_{t-1} + u_t$$

$$\Delta y_t = \gamma y_{t-1} + u_t$$

- $H_0 : \gamma = 0$  indicates a unit root.
  - For stationarity, we reject in favor of  $\gamma < 0$ . (note this is a one-sided test)
  - In this simple form, this test is known as the "Dickey-Fuller Test".
  - Test statistics are not based on a t-distribution - Table in book, correct p-values given in R.

# What is a unit-root? (graphically)

- Create 3 different AR(1) time series, at or near a unit root.

```
Nobs<-100
```

```
x<-AR1(Nobs,0.25)
```

```
y<-AR1(Nobs,1)
```

```
z<-AR1(Nobs,1.02)
```

- Plot the time series

```
par(mfrow=c(1,3))
```

```
plot(x,type='l',main="phi=0.25",xlab='t',ylab="Y")
```

```
plot(y,type='l',main="phi=1",xlab='t',ylab="Y")
```

```
plot(z,type='l',main="phi=1.02",xlab='t',ylab="Y")
```

- What are the features of these three plots?

# Integrated series

- From the previous series...

$$\Delta y_t = \gamma y_{t-1} + u_t$$

- Again, this is stationary when  $\gamma < 0$ .
  - This type of series is called "integrated of order 0": I(0)
- If not stationary, take differences and test again. If  $\Delta y_t$  is stationary, this type of series is called "integrated of order 1"
- In general, a series is integrated order  $d$  if  $d$  differences are required to make stationary.
- In R, for our previous series, take differences and plot:

```
plot(diff(x, lag=1), type='l', main="phi=0.25", xlab='t', ylab="Y")
plot(diff(y, lag=1), type='l', main="phi=1", xlab='t', ylab="Y")
plot(diff(z, lag=1), type='l', main="phi=1.02", xlab='t', ylab="Y")
```
- Do the differenced series look more stationary?

# Testing for unit roots manually

- In R, we need to regress the differences of a time series on initial values and test the coefficient.

- Using our original time series  $x$ :

```
summary(lm(diff(x, lag=1) ~ x[1:(N-1)]))
```

- $x[1:(N-1)]$  is the vector of matched initial time periods.

- Do the same for the other series

```
summary(lm(diff(y, lag=1) ~ y[1:(N-1)]))
```

```
summary(lm(diff(z, lag=1) ~ z[1:(N-1)]))
```

- These regressions are only suggestive in significance. Must use the DF significance table from the book, or the R code that I will present in a few slides.

# Testing for unit roots manually

- When series appear to be non-stationary, we need find out how many differences we need to take for it to be stationary.
- To begin, test for stationarity in the differenced data
- Formally, we are testing where  $\phi$  lies relative to 1 in the following

$$\Delta y_t = \phi \Delta y_{t-1} + u_t$$

- Subtracting  $\Delta y_{t-1}$  from both sides

$$\begin{aligned}\Delta y_t - \Delta y_{t-1} &= \phi \Delta y_{t-1} - \Delta y_{t-1} + u_t \\ \Delta^2 y_t &= \gamma \Delta y_{t-1} + u_t\end{aligned}$$

- In R, for some series  $z$ :

```
d1<-z [2 : (N-1) ] -z [1 : (N-2) ]  
d2<-z [3 :N] -z [2 : (N-1) ]  
summary (lm (I (d2-d1) ~d1))
```

# Stationarity tests in R

- In the package "tseries", `adf.test(x,k=0)` runs the standard DF test.

```
library(tseries)
adf.test(x, k=0)
adf.test(y, k=0)
adf.test(z, k=0)
```

- The null hypothesis is that there is a unit root, and the alternative is "stationary".
- The augmented DF test runs the following regression

$$\Delta y_t = \beta_0 + \alpha t + \gamma y_{t-1} + \sum_{i=1}^k \beta_i \Delta y_{t-i} + e_t$$

- $k$  adjusts the lag length in the regression.

```
adf.test(z, k=0)
adf.test(z, k=1)
adf.test(z, k=2)
```

- Again, the null is that there is a unit root.



# Why we care - Spurious Regressions

- Recall from our earlier example that when  $\phi = 1$

$$y_t = y_{t-1} + u_{yt}$$

- If we run this process enough times what do we notice?
  - The series usually trends somewhere.

- Suppose we have an independently constructed series of the same form:

$$x_t = x_{t-1} + u_{xt}$$

- If we regress  $y_t$  on  $x_t$ , what happens?
- Since both series have a tendency to trend somewhere, there will appear to be a relationship between the two series most of the time.
- This is called a **spurious relationship**. We must therefore identify unit roots when regressing time series on one another to prevent this issue.

# Why we care - Spurious Regressions

- Regressing two independently created series should not show a systematic relationship
- But, when there is a unit root in both series, there may be a spurious relationship.
  - This is bad for macro data, since as we know aggregate variables usually trend somewhere.
- Run this code repeatedly and see how many times you get an insignificant relationship between the two series

```
x1<-AR1(N,1)
x2<-AR1(N,1)
summary(lm(x2~x1))
```

- Along with there clearly being no mechanical relationship between the series (since they are random), the standard errors are incorrect for classic OLS

# Spurious Regressions (cont.)

- Why are standard errors incorrect?
- Suppose we wish to regress  $y_t$  on  $x_t$  using

$$y_t = \beta_0 + \beta_1 x_t + u_t$$

- Rearranging for  $u_t$ , we have:

$$u_t = y_t - \beta_0 - \beta_1 x_t$$

- Back-substituting for  $y_t = y_{t-1} + u_{yt-1}$  and  $x_t = x_{t-1} + u_{xt-1}$ , we have:

$$u_t = (y_{t-1} + u_{yt-1}) - \beta_0 - \beta_1 (x_{t-1} + u_{xt-1})$$

- Doing so repeatedly for back to period 1, we get:

$$u_t = y_1 + \sum_{i=1}^{t-1} u_{yi} + \beta_0 - \beta_1 x_1 - \beta_1 \sum_{i=1}^{t-1} u_{xi}$$

- Note that because of the  $\sum_{i=1}^{t-1} u_{yi}$  and  $\beta_1 \sum_{i=1}^{t-1} u_{xi}$  the variance explodes as  $t$  gets large relative to the initial state.

# Cointegration and Error-Correction

- Trending time series cause problems due to the spurious regression
  - This tends to be a problem with any macro data
- Differencing helps, but there are drawbacks
  - Cannot speak to long-run changes, only short-run (since identifying variation is based on first differences or higher order differences)
- **Cointegration** provides a framework for identifying and estimating time series regressions

# Definition of Cointegration

- Suppose we have the following time-series model

$$Y_t = \beta_0 + \beta_1 X_t + u_t$$

- Estimate the model to get  $\hat{\beta}_0$  and  $\hat{\beta}_1$ . Constructing the residuals:

$$\hat{u}_t = Y_t - \hat{\beta}_0 - \hat{\beta}_1 X_t$$

- If  $\hat{u}_t \sim I(0)$ , then  $Y_t$  and  $X_t$  are **cointegrated**
- This is trivial if  $Y_t$  and  $X_t$  are both  $I(0)$
- This more interesting when  $Y_t$  and  $X_t$  are both  $I(1)$
- How often does this occur given random time series generated by a process with a unit-root?

# Cointegrated series - Monte Carlo examples

- Insert examples, theory

```
for(i in 1:1000) {  
  x<-AR1(N,1)  
  y<-AR1(N,1)  
  errors<-resid(lm(y~x))  
  adftest<-adf.test(errors)  
  p<-adftest$p.value  
  if(i==1) {res<-data.frame(p)}  
  if(i>1) {res<-rbind(res,data.frame(p))}  
}
```

- Calculate how many reject a unit root in favor of stationarity

```
mean(res$p<0.1, na.rm=TRUE)
```

# Error Correction Model

- If  $Y_t$  and  $X_t$  are  $I(1)$ , but  $\hat{u}_t \sim I(0)$ , then we can estimate using OLS the following "Error correction model"

$$\Delta Y_t = a_0 + b_1 \Delta X_t + \pi \hat{u}_{t-1} + e_t$$

- This regression is **not** spurious because  $\hat{u}_{t-1}$ ,  $\Delta Y_t$  and  $\Delta X_t$  are all  $I(0)$
- Since,  $Y_t$  and  $X_t$  are also  $I(0)$ , we can obtain consistent estimates for  $b_1$  using standard regression
- We can also obtain long-run equilibrium dynamics by focusing on  $\pi$ .
  - $\hat{u}_{t-1} \neq 0$  indicates *disequilibrium* between  $Y$  and  $X$  in  $Y_{t-1} = \beta_0 + \beta_1 X_{t-1} + u_{t-1}$
  - $\pi = -1$  equilibrium is reached immediately
  - $\pi \in (-1, 0]$  equilibrium is reached gradually
  - $\pi < -1$  suggests an over-correction

# Error Correction Model to ARDL model

- The ECM model is equivalent to a ARDL model (Autoregressive Distributed Lag), which we presented a few lectures ago when talking about VARs. To see this, note that:

$$\Delta Y_t = a_0 + b_1 \Delta X_t + \pi \hat{u}_{t-1} + e_t$$

- Expanding  $\Delta Y_t$ ,  $\Delta X_t$ , and  $\hat{u}_{t-1}$ , we have:

$$Y_t - Y_{t-1} = a_0 + b_1 (X_t - X_{t-1}) + \pi (Y_{t-1} - \beta_0 - \beta_1 X_{t-1}) + e_t$$

- Bringing all lags to the RHS:

$$Y_t = a_0 + Y_{t-1} + \pi Y_{t-1} + b_1 X_t - b_1 X_{t-1} - \pi \beta_0 - \pi \beta_1 X_{t-1} + e_t$$

- Collecting terms

$$Y_t = a_0 - \pi \beta_0 + (1 + \pi) Y_{t-1} + b_1 X_t + (-b_1 - \pi \beta_1) X_{t-1} + e_t$$

- Thus, we have an ARDL model.



# Engel-Granger Technique

- Engel and Granger have proposed a technique for evaluating data that may be spurious.
- **Step 1:** Determine whether  $X_t$  and  $Y_t$  are cointegrated.
  - If  $X_t$  and  $Y_t$  are  $I(0)$ , then use classic regression
  - If only one of  $X_t$  and  $Y_t$  are  $I(1)$ , and the other  $I(0)$ , then need a new technique
  - If  $X_t$  and  $Y_t$  are  $I(1)$ , then run

$$Y_t = \beta_0 + \beta_1 X_t + u_t$$

and collect residuals. Go to step 2.

- **Step 2:** Check whether  $u_t$  is  $I(0)$ 
  - If  $u_t$  is  $I(0)$ , move to step 3.
  - If  $u_t$  is  $I(1)$ , find a new model
- **Step 3:** Estimate and interpret:

$$\Delta Y_t = a_0 + b_1 \Delta X_t + \pi \hat{u}_{t-1} + e_t$$

# Example - Co-integration of investment funds

- Download a year of daily opening prices for SPY and VOO

```
getSymbols('SPY', from='2014-11-12', to='2015-11-12')
```

```
getSymbols('VOO', from='2014-11-12', to='2015-11-12')
```

```
prices.spy <- SPY$SPY.Open
```

```
prices.voo <- VOO$VOO.Open
```

- **Step 1:** Determine whether *SPY* and *VOO* have unit root.

```
adf.test(prices.spy)
```

```
adf.test(prices.voo)
```

# Example - Co-integration of investment funds

- **Step 2:** Regress

$$SPY_t = a_0 + b_1 VOO_t + u_t$$

and test whether  $u_t$  is  $I(0)$

```
coint <- lm(prices.spy~prices.voo)
summary(coint)
beta<-coint$coef
resid <- prices.spy - (beta[1] + beta[2]*prices.voo)
adf.test(resid)
```

- **Step 3:** Estimate and interpret:

$$\Delta SPY_t = a_0 + b_1 \Delta VOO_t + \pi \hat{u}_{t-1} + e_t$$

```
dSPY<-prices.spy-lag(prices.spy,1)
dVOO<-prices.voo-lag(prices.voo,1)
ecm <-lm(dSPY~dVOO+lag.resid)
summary(ecm)
```