

# Bayesian Inference (I)

Intro to Bayesian Data Analysis & Cognitive Modeling  
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[based on slides by Sharon Goldwater & Frank Keller]

Fall 2012 · UCSC Linguistics

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How do people make decisions? For example,

- Medicine: Which disease to diagnose?
- Business: Where to invest? Whom to trust?
- Law: Whether to convict?
- Admissions/hiring: Who to accept?
- Language interpretation: What meaning to select for a word? How to resolve a pronoun? What quantifier scope to choose for a sentence?

In all these cases, we use two kinds of information:

- Background knowledge:
  - prevalence of disease
  - previous experience with business partner
  - historical rates of return in market
  - relative frequency of the meanings of a word
  - scoping preference of a quantifier
  - etc.
- Specific information about this case:
  - test results
  - facial expressions and tone of voice
  - company business reports
  - various features of the current sentential and discourse context
  - etc.

Example question from a study of decision-making for medical diagnosis (Casscells et al. 1978):

## Example

If a test to detect a disease whose prevalence is 1/1000 has a false-positive rate of 5%, what is the chance that a person found to have a positive result actually has the disease, assuming you know nothing about the person's symptoms or signs?

*Most frequent answer: 95%*

Reasoning: if false-positive rate is 5%, then test will be correct 95% of the time.

*Correct answer: about 2%*

Reasoning: assume you test 1000 people; only about one person actually has the disease, but the test will be positive in another 50 or so cases (5%). Hence the chance that a person with a positive result has the disease is about  $1/50 = 2\%$ .

Only 12% of subjects give the correct answer.

*Mathematics underlying the correct answer: Bayes' Theorem.*

To analyze the answers that subjects give, we need:

### Bayes' Theorem

Given a hypothesis  $h$  and data  $D$  which bears on the hypothesis:

$$p(h|D) = \frac{p(D|h)p(h)}{p(D)}$$

$p(h)$ : independent probability of  $h$ : *prior probability*

$p(D)$ : independent probability of  $D$ : *marginal likelihood / evidence*

$p(D|h)$ : conditional probability of  $D$  given  $h$ : *likelihood*

$p(h|D)$ : conditional probability of  $h$  given  $D$ : *posterior probability*

We also need the *rule of total probability*.

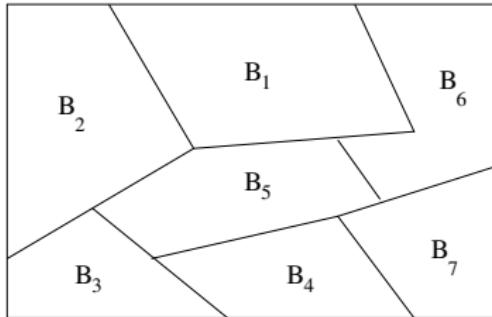
### Theorem: Rule of Total Probability

If events  $B_1, B_2, \dots, B_k$  constitute a partition of the sample space  $S$  and  $p(B_i) \neq 0$  for  $i = 1, 2, \dots, k$ , then for any event  $A$  in  $S$ :

$$p(A) = \sum_{i=1}^k p(A|B_i)p(B_i)$$

$B_1, B_2, \dots, B_k$  form a **partition** of  $S$  if they are pairwise mutually exclusive and if

$$B_1 \cup B_2 \cup \dots \cup B_k = S.$$



# Evidence/Marginal Likelihood and Bayes' Theorem

## Evidence/Marginal Likelihood

The **evidence** is also called the **marginal likelihood** because it is the likelihood  $p(D|h)$  marginalized relative to the prior probability distribution over hypotheses  $p(h)$ :

$$p(D) = \sum_h p(D|h)p(h)$$

It is also sometimes called the **prior predictive distribution** because it provides the average/mean probability of the data  $D$  given the prior probability over hypotheses  $p(h)$ .

## Reexpressing Bayes' Theorem

Given the above formula for the evidence, Bayes' theorem can be alternatively expressed as:

$$p(h|D) = \frac{p(D|h)p(h)}{\sum_h p(D|h)p(h)}$$

# Bayes' Theorem for Data $D$ and Model Parameters $\theta$

In the specific case of a model with parameters  $\theta$  (e.g., the bias of a coin), Bayes' theorem is:

$$p(\theta_j|D_i) = \frac{p(D_i|\theta_j)p(\theta_j)}{\sum_{j \in J} p(D_i|\theta_j)p(\theta_j)}$$

		parameter values			
data values		$\theta_j$			
		...	...	...	
	$D_i$	...	$p(D_i, \theta_j)$ $= p(D_i \theta_j)p(\theta_j)$ $= p(\theta_j D_i)p(D_i)$	...	
		...	...	...	$p(D_i) = \sum_{j \in J} p(D_i \theta_j)p(\theta_j)$
		...	...	...	...
		$p(\theta_j)$			

## Application of Bayes' Theorem

In Casscells et al.'s (1978) example, we have:

- $h$ : person tested has the disease;
- $\bar{h}$ : person tested doesn't have the disease;
- $D$ : person tests positive for the disease.

$$p(h) = 1/1000 = 0.001 \quad p(\bar{h}) = 1 - p(h) = 0.999$$

$$p(D|\bar{h}) = 5\% = 0.05 \quad p(D|h) = 1 \text{ (assume perfect test)}$$

Compute the probability of the data (rule of total probability):

$$p(D) = p(D|h)p(h) + p(D|\bar{h})p(\bar{h}) = 1 \cdot 0.001 + 0.05 \cdot 0.999 = 0.05095$$

Compute the probability of correctly detecting the illness:

$$p(h|D) = \frac{p(h)p(D|h)}{p(D)} = \frac{0.001 \cdot 1}{0.05095} = 0.01963$$

**Base rate:** the probability of the hypothesis being true in the absence of any data, i.e.,  $p(h)$  (the prior probability of disease).

**Base rate neglect:** people tend to ignore / discount base rate information, as in Casscells et al.'s (1978) experiments.

- has been demonstrated in a number of experimental situations;
- often presented as a fundamental bias in decision making.

Does this mean people are irrational/sub-optimal?

Casscells et al.'s (1978) study is abstract and artificial. Other studies show that

- data presentation affects performance (1 in 20 vs. 5%);
- direct experience of statistics (through exposure to many outcomes) affects performance;  
(which is why you should tweak the R and JAGS code in this class extensively and try it against a lot of simulated data sets)
- task description affects performance.

Suggests subjects may be interpreting questions and determining priors in ways other than experimenters assume.

Evidence that subjects can use base rates: diagnosis task of Medin and Edelson (1988).

Bayesian interpretation of probabilities is that they reflect *degrees of belief*, not frequencies.

- Belief can be influenced by frequencies: observing many outcomes changes one's belief about future outcomes.
- Belief can be influenced by other factors: structural assumptions, knowledge of similar cases, complexity of hypotheses, etc.
- Hypotheses can be assigned probabilities.

# Bayes' Theorem, Again

$$p(h|D) = \frac{p(D|h)p(h)}{p(D)}$$

$p(h)$ : *prior probability* reflects plausibility of  $h$  regardless of data.

$p(D|h)$ : *likelihood* reflects how well  $h$  explains the data.

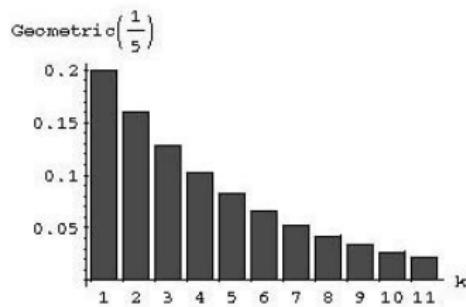
$p(h|D)$ : *posterior probability* reflects plausibility of  $h$  after taking data into account.

Upshot:

- $p(h)$  may differ from the “base rate” / counting
- the base rate neglect in the early experimental studies might be due to equating probabilities with relative frequencies
- subjects may use additional information to determine prior probabilities (e.g., if they are wired to do this)

So far, we have discussed *discrete distributions*.

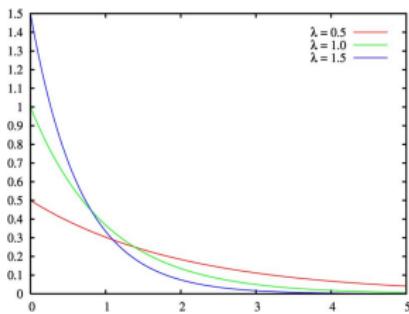
- Sample space  $S$  is finite or countably infinite (integers).
- Distribution is a *probability mass function*, defines probability of r.v. having a particular value.
- Ex:  $p(Y = n) = (1 - \theta)^{n-1} \theta$  (Geometric distribution):



(Image from <http://eom.springer.de/G/g044230.htm>)

We will also see *continuous distributions*.

- Support is uncountably infinite (real numbers).
- Distribution is a *probability density function*, defines relative probabilities of different values (sort of).
- Ex:  $p(Y = y) = \lambda e^{-\lambda y}$  (Exponential distribution):



(Image from Wikipedia)

Discrete distributions ( $p(\cdot)$  is a probability mass function):

- $0 \leq p(Y = y) \leq 1$  for all  $y \in S$
- $\sum_y p(Y = y) = \sum_y p(y) = 1$
- $p(y) = \sum_x p(y|x)p(x)$  (Law of Total Prob.)
- $E[Y] = \sum_y y \cdot p(y)$  (Expectation)

Continuous distributions ( $p(\cdot)$  is a probability density function):

- $p(y) \geq 0$  for all  $y$
- $\int_{-\infty}^{\infty} p(y)dy = 1$  (if the support of the dist. is  $\mathbb{R}$ )
- $p(y) = \int_x p(y|x)p(x)dx$  (Law of Total Prob.)
- $E[X] = \int_x x \cdot p(x)dx$  (Expectation)

Simple inference task: estimate the probability that a particular coin shows heads. Let

- $\theta$ : the probability we are estimating.
- $H$ : hypothesis space (values of  $\theta$  between 0 and 1).
- $D$ : observed data (previous coin flips).
- $n_h, n_t$ : number of heads and tails in  $D$ .

Bayes' Rule tells us:

$$p(\theta|D) = \frac{p(D|\theta)p(\theta)}{p(D)} \propto p(D|\theta)p(\theta)$$

How can we use this for predictions?

# Maximum Likelihood Estimation

1. Choose  $\theta$  that makes  $D$  most probable, i.e., ignore  $p(\theta)$ :

$$\hat{\theta} = \operatorname{argmax}_{\theta} p(D|\theta)$$

This is the *maximum likelihood* (ML) estimate of  $\theta$ , and turns out to be equivalent to relative frequencies (proportion of heads out of total number of coin flips):

$$\hat{\theta} = \frac{n_h}{n_h + n_t}$$

- Insensitive to sample size (10 coin flips vs 1000 coin flips), and does not generalize well (overfits).

# Maximum A Posteriori Estimation

2. Choose  $\theta$  that is most probable given  $D$ :

$$\hat{\theta} = \operatorname{argmax}_{\theta} p(\theta|D) = \operatorname{argmax}_{\theta} p(D|\theta)p(\theta)$$

This is the *maximum a posteriori* (MAP) estimate of  $\theta$ , and is equivalent to ML when  $p(\theta)$  is uniform.

- Non-uniform priors can reduce overfitting, but MAP still doesn't account for the shape of  $p(\theta|D)$ :



# Posterior Distribution and Bayesian Integration

3. Work with the entire posterior distribution  $p(\theta|D)$ .

Good measure of central tendency – the expected posterior value of  $\theta$  instead of its maximal value:

$$E[\theta] = \int \theta p(\theta|D) d\theta = \int \theta \frac{p(D|\theta)p(\theta)}{p(D)} d\theta \propto \int \theta p(D|\theta)p(\theta) d\theta$$

This is the *posterior mean*, an average over hypotheses. When prior is uniform (i.e.,  $Beta(1, 1)$ , as we will soon see), we have:

$$E[\theta] = \frac{n_h + 1}{n_h + n_t + 2}$$

- Automatic smoothing effect: unseen events have non-zero probability.

Anything else can be obtained out of the posterior distribution: median, 2.5% and 97.5% quantiles, any function of  $\theta$  etc.

## E.g.: Predictions based on MAP vs. Posterior Mean

Suppose we need to classify inputs  $y$  as either positive or negative, e.g., indefinites as taking wide or narrow scope.

There are only 3 possible hypotheses about the correct method of classification (3 theories of scope preference):  $h_1$ ,  $h_2$  and  $h_3$  with posterior probabilities 0.4, 0.3 and 0.3, respectively.

We are given a new indefinite  $y$ , which  $h_1$  classifies as positive / wide scope and  $h_2$  and  $h_3$  classify as negative / narrow scope.

- using the MAP estimate, i.e., hypothesis  $h_1$ ,  $y$  is classified as wide scope
- using the posterior mean, we average over all hypotheses and classify  $y$  as narrow scope

- Casscells, W., A. Schoenberger, and T. Grayboys: 1978, 'Interpretation by Physicians of Clinical Laboratory Results', *New England Journal of Medicine* **299**, 999–1001.
- Medin, D. L. and S. M. Edelson: 1988, 'Problem Structure and the Use of Base-rate Information from Experience', *Journal of Experimental Psychology: General* **117**, 68–85.