# HW2 Answer Key 

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## Problem 1

(a) The log-likelihood is

$$
l=\sum_{i=1}^{n}\left[-\frac{1}{2} \log \left(2 \pi \sigma^{2}\right)-\frac{1}{2 \sigma^{2}}\left(y_{i}-\mu_{i}\right)^{2}\right]=-\frac{n}{2} \log \left(2 \pi \sigma^{2}\right)-\frac{1}{2 \sigma^{2}} \sum_{i=1}^{n}\left(y_{i}-\mu_{i}\right)^{2}
$$

(b) Given $\mu_{i}=\mu$, we choose $\hat{\mu}$ to maximize the log-likelihood, i.e.,

$$
\max _{\hat{\mu}} l=-\frac{n}{2} \log \left(2 \pi \sigma^{2}\right)-\frac{1}{2 \sigma^{2}} \sum_{i=1}^{n}\left(y_{i}-\hat{\mu}\right)^{2}
$$

First-order condition implies that

$$
\frac{1}{\sigma^{2}} \sum_{i=1}^{n}\left(y_{i}-\hat{\mu}\right)=0
$$

which simplifies to

$$
n \hat{\mu}-\sum_{i=1}^{n} y_{i}=0
$$

Therefore, the maximum likelihood estimate for $\mu$ is

$$
\hat{\mu}=\frac{1}{n} \sum_{i=1}^{n} y_{i}
$$

i.e., the arithmetic mean of outcomes.

## Problem 2

(a) First, note that $f^{\prime}=3\left(e^{x}-2\right)^{2} e^{x}$, therefore, we use

$$
x_{i+1}=x_{i}-\frac{f\left(x_{i}\right)}{f^{\prime}\left(x_{i}\right)}=x_{i}-\frac{\left(e^{x_{i}}-2\right)^{3}}{3\left(e^{x_{i}}-2\right)^{2} e^{x_{i}}}=x_{i}-\frac{e^{x_{i}}-2}{3 e^{x_{i}}}=x_{i}+\frac{2}{3} e^{-x_{i}}-\frac{1}{3}
$$

to iterate for solution. ${ }^{1}$ The code for function is:

```
NR<-function(x){
    counter <- 1
    while( abs(-1/3+2*(exp(-x))/3)>1e-10 & counter<=300000 ){
        x <- x-1/3+2* (exp(-x))/3
        counter <- counter+1
    }
    x
}
```

Note that, the condition abs $(-1 / 3+2 *(\exp (-\mathrm{x})) / 3)>1 \mathrm{e}-10$ represents $\left|x_{i+1}-x_{i}\right|>0$, which indicates that an additional iteration does not refine the value of $x$.

Then, simply by typing NR(0) we will get the answer 0.6931472 .

[^0](b) Now, $f(x)=\left(e^{x}-2\right)^{n}$, so $f^{\prime}=n\left(e^{x}-2\right)^{n-1} e^{x}$. Similar to Part (a), the NR iteration is
$$
x_{i+1}=x_{i}-\frac{f\left(x_{i}\right)}{f^{\prime}\left(x_{i}\right)}=x_{i}-\frac{\left(e^{x_{i}}-2\right)^{n}}{n\left(e^{x_{i}}-2\right)^{n-1} e^{x_{i}}}=x_{i}-\frac{e^{x_{i}}-2}{3 e^{x_{i}}}=x_{i}+\frac{2}{n} e^{-x_{i}}-\frac{1}{n}
$$

We modify the interation function in R so that $n$ enters as an argument. Further, the returning value now (c(x, counter)) is a vector that includes both the solution and the number of iteration.

```
NR<-function(x,n){
    counter <- 1
    while( abs(-1/n+2*(exp(-x))/n)>1e-10 & counter<=30000){
        x <- x-1/n+2* (exp (-x))/n
        counter <- counter+1
    }
    c(x,counter)
}
```

By entering from $\operatorname{NR}(0,2)$ to $\operatorname{NR}(0,8)$, you will see

| $n$ | Solution | Number of Iterations |
| :---: | :---: | :---: |
| 2 | 0.6931472 | 32 |
| 3 | 0.6931472 | 53 |
| 4 | 0.6931472 | 74 |
| 5 | 0.6931472 | 94 |
| 6 | 0.6931472 | 114 |
| 7 | 0.6931472 | 133 |
| 8 | 0.6931472 | 153 |

Notes: depending on your criteria, you may get different number of iterations.

## Problem 3

(a) We use Poisson regression to specify the relation between real wage and our attributes of interest. By running the following code

```
library(foreign)
d<-read.dta("D:/org_example.dta")
ds<-subset(d,state=="CA" & year==2013)
poissonreg<-glm(rw~educ+female+age+wbho,ds,family="poisson"(link="log"))
summary(poissonreg)
```

You will get the result
Call:
glm(formula $=r w \sim$ educ + female + age + wbho, family $=$ poisson(link $=$ "log"),

```
    data = ds)
```

Deviance Residuals:

| Min | 1Q | Median | 3Q | Max |
| ---: | ---: | ---: | ---: | ---: |
| -9.9092 | -1.8621 | -0.6145 | 0.9679 | 25.9487 |

Coefficients:

|  | Estimate Std. Error z value $\operatorname{Pr}(>\|\mathrm{z}\|)$ |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 2.2051037 | 0.0151662 | 145.40 | $<2 \mathrm{e}-16 * * *$ |  |
| educHS | 0.3397354 | 0.0125034 | 27.17 | $<2 \mathrm{e}-16 * * *$ |  |


| educSome college | 0.4720058 | 0.0122888 | 38.41 | <2e-16 | *** |
| :---: | :---: | :---: | :---: | :---: | :---: |
| educCollege | 0.9353610 | 0.0123980 | 75.44 | $<2 \mathrm{e}-16$ | *** |
| educAdvanced | 1.1797471 | 0.0126980 | 92.91 | $<2 \mathrm{e}-16$ | *** |
| female | -0.1918608 | 0.0050182 | -38.23 | $<2 \mathrm{e}-16$ | *** |
| age | 0.0111549 | 0.0001857 | 60.08 | $<2 \mathrm{e}-16$ | * |
| wbhoBlack | -0.2713582 | 0.0129974 | -20.88 | $<2 \mathrm{e}-16$ | *** |
| wbhoHispanic | -0.1710364 | 0.0067647 | -25.28 | $<2 \mathrm{e}-16$ | * |
| wbhoOther | -0.0875031 | 0.0066407 | -13.18 | $<2 \mathrm{e}-16$ | *** |
|  |  |  |  |  |  |
| Signif. codes: 0 *** $0.001 * * 0.01 * 0.05 .0 .1 \quad 1$ |  |  |  |  |  |
| (Dispersion parameter for poisson family taken to be 1) |  |  |  |  |  |
| Null deviance: 83415 on 6840 degrees of freedom |  |  |  |  |  |
| Residual deviance: 50183 on 6831 degrees of freedom (6948 observations deleted due to missingness) |  |  |  |  |  |
| AIC: Inf |  |  |  |  |  |

Given other variables fixed, a female earns $19.2 \%$ less than a male, given other factors fixed.
Notes: some students replace NA in outcomes with 0 and get a different result (around 30\%), or directly calculate $\mathrm{E}[Y \mid$ female $=1]-\mathrm{E}[Y \mid$ female $=0]$ rather than the marginal effect. They are also correct.
(b) Here is the LM test. Note that the variable wbho includes four choices, thus the degree of freedom is 3 .

```
poissonreg2<-glm(rw~ educ+female+age,ds,family="poisson"(link="log"))
LR<-(poissonreg2$deviance-poissonreg$deviance)
chi_crit<-qchisq(.95, df=3)
ifelse(LR>chi_crit,"Reject the restrictions", "Fail to reject the restrictions")
```

The null hypothesis is rejected at $95 \%$ significance level, i.e., ethnicity does play a role in wage outcomes.
(c) Using the regression from part (a), we use the following code:

```
ds$glm_predict<-predict(poissonreg,newdata=ds,type="response")
plot(density(ds$rw,na.rm=TRUE),lwd=1.5,main="Real wages (black) vs Predicted Real Wages (blue)",
    xlab="Real Wages")
lines(density(ds$glm_predict,na.rm=TRUE),lwd=2,col="blue")
```

The output is in the figure below. Here, we see that while the distributions of actual values and predicted values are pretty similar, the mode of the predicted values is skewed to the right a bit. This is because there are some very high wages in the data, which pull up the average from zero. However, the independent variables do not have sufficient variation to explain these high values, so the regression compensates by increasing the predicted value of other observations so that the means are the same for both predicted and actual values.

Real wages (black) vs Predicted Real Wages (blue)


## Problem 4

(a)
d<-read.dta("D:/org_example.dta")
ds<-subset (d,state=="CA" \& year==2013)
ds\$logrw<-log (ds\$rw)
ds\$loghourslw<-log(ds\$hourslw)
fit.loess1 <- loess(dsfull\$logrw~dsfull\$hourslw,span = 1,degree= 2)
x1=seq(min(dsfull\$hourslw), max (dsfull\$hourslw), 0.01)
plot(x1,predict(fit.loess1, x1), col = "black", xlab="Hours worked",ylab="Log real wage")
fit.loess2 <- loess(dsfull\$logrw~dsfull\$loghourslw,span = 1, degree= 2)
x2=seq(min(dsfull\$loghourslw), max (dsfull\$loghourslw), 0.01)
plot(x2,predict(fit.loess1, x2), col = "black", xlab="Log hours worked",ylab="Log real wage")


The
relation between hours worked and log wage exhibits bell shape: when working hours is less than about 65 per week, the log hourly wage increases with working hours, but when working hours is over 65 per week, the $\log$ hourly wage decreases with working hours. However, if we estimate the relation between log hours worked and $\log$ wage, we find a positive and proportional relation.
(b) Using the regression from part (a), we use the following code:

```
gamresults<-gam(ds$logrw ~s(ds$hourslw)+educ+age, data=ds)
summary(gamresults)
plot(gamresults,se=TRUE,rug=FALSE,terms="s")
abline(h=0)
```

We find that when the working hours is between 5 and 35 , the predicted $\log$ wage is significantly lower than average, and when the working hours is between 40 and 70 , the predicted $\log$ wage is significantly higher than average.

(c)

```
small<- data.frame(dsfull$logrw,dsfull$hourslw)
for(h in 1:20){
    for(i in 1:nrow(small)){
        smalldrop<-small[i,]
        smallkeep<-small[-i,]
```

```
    fit<-loess(small$logrw~}small$hourslw,smallkeep, family="gaussian",span=(h/20), degree=1
    dropfit<-predict(fit,smalldrop,se=FALSE)
    sqrerr<-(smalldrop$logrw-as.numeric(dropfit))^2
    if(i*h==1){results<-data.frame(h,i,sqrerr)}
    if(i*h>1){results<-rbind(results,data.frame(h,i,sqrerr))}
} }
tapply(results$sqrerr,results$h,FUN=sum,na.rm=TRUE)
```

We find the optimal span is 0.65 in this case.


[^0]:    ${ }^{1}$ Of course you may use the original expression. However, using simplified equation will accelerate the computation and make it more precise.

